

# Downward Rigidity in the Wage for New Hires

Joe Hazell (LSE)

Bledi Taska (Burning Glass Technologies)

Yale

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# Importance of Wage Rigidity for New Hires

- If wages are rigid downward, unemployment rises sharply during recessions  
(Keynes 1936, Hall 2005)
- **Wage for new hires** is particularly important (Pissarides 2009)
- **Limited evidence** of downward rigidity for new hires

# Job Composition: Challenge for Measuring New Hire Wage

- Previous work: **average** wage for new hires (e.g. Haefke, Sonntag & van Rens 2013)
- Key challenge: **job composition** (Gertler & Trigari 2009)

$$\Delta \text{average wage} = \Delta \text{job-level wage} + \Delta \text{job composition}$$

- E.g.: high wage bankers + low wage baristas
  - Share of barista hires rises
  - Average wage for hires falls even if wages do not fall for bankers + baristas
- Previous estimates are **imprecise**

# Job-Level Data on the Wage for New Hires

- This paper: online dataset of **wage for new hires** (Burning Glass)
  - Establishment level wages **posted on vacancies** + job titles
  - 15% of total US vacancies during 2010-2021
- Burning Glass tracks other measures of average wage for new hires

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  - Establishment level wages **posted on vacancies** + job titles
  - 15% of total US vacancies during 2010-2021
- Burning Glass tracks other measures of average wage for new hires
- We can measure **job-level** wage:
  - Between successive vacancies of job title and establishment
- Example: a physical establishment of Starbucks
  - Posts vacancies for baristas, pays an hourly wage

# Wage for New Hires: Rigid Downward But Flexible Upward

Wage for new hires is **rigid downward but flexible upward**

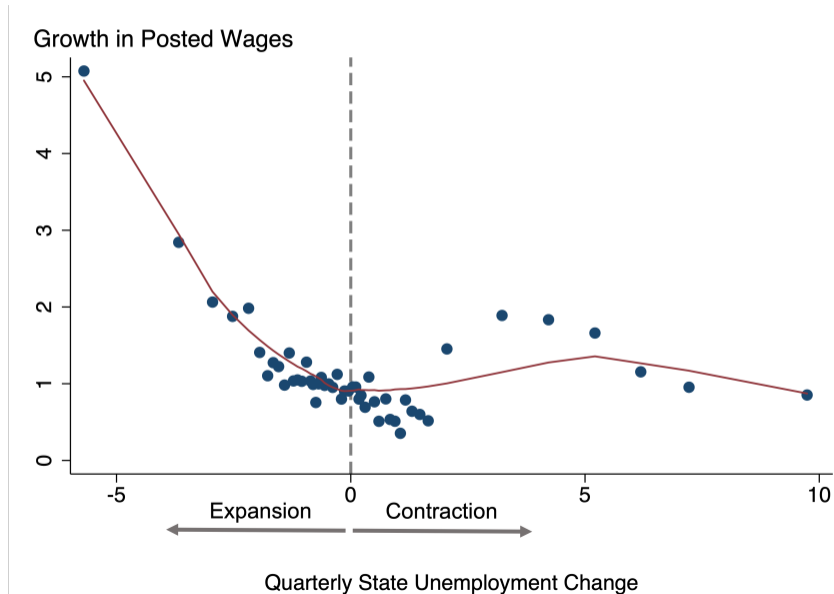
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2. Job level wages respond to expansions, do not respond to contractions
3. No downward rigidity in average wage due to job composition
4. Downward wage rigidity → state dependent wage flexibility upward

# Wage for New Hires: Rigid Downward But Flexible Upward

Wage for new hires is **rigid downward but flexible upward**

1. Job level wages rarely change, rise more often than fall
  - Duration of unchanged wages is 3.3 quarters
  - Conditional on change, increase is  $\sim 3$  times more likely
2. Job level wages respond to expansions, do not respond to contractions
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# Literature

- **Shimer puzzle and wage rigidity** (Shimer 2005; Hall 2005; Hall + Milgrom 2008; Gertler & Trigari 2009; Christiano, Eichenbaum & Trabandt 2016, 2021)  
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- **Worker-level data on wage rigidity for new hires** (Bils 1985; Haefke, Sonntag & van Rens 2013; Gertler, Trigari & Huckfeldt 2020; Grigsby, Hurst & Yildirmaz 2021)  
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Contribution: (i) job-level data (ii) downward constraint (iii) “hockey stick”
- **Downward rigidity and asymmetric unemployment dynamics**  
(Schmitt-Grohé & Uribe 2016; Chodorow-Reich & Wieland 2020; Dupraz, Nakamura & Steinsson 2021; Guerrieri, Lorenzoni, Straub & Werning 2021)  
Contribution: new evidence on asymmetry

# Dataset: Posted Wages in Burning Glass

## Burning Glass Data (1/2)

Online vacancies from Burning Glass Technologies:

- Establishment level dataset of vacancies, with job titles, industry + occ info
- Reports posted wage with pay frequency and bonus + overtime pay
- $\sim 1/3$  of vacancies post a range, the rest post point wages
- Time period: 2010-2021Q1

Sample selection:

- Online vacancies are  $\sim 80\%$  of total US vacancies
  - Only 20% of online vacancies post wages
- Our sample (online vacancies with wages)  $\sim 15\%$  of total US vacancies



# Burning Glass Data (2/2)

Potential concerns with data:

#1 Data overweights certain occupations [Details](#)

Solution: occupation reweighting, dataset has granular occupation coverage

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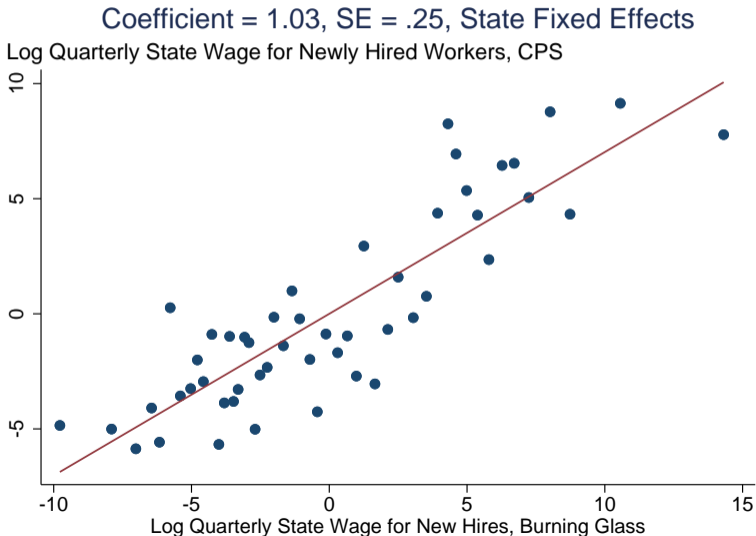
Solution: selection is uncorrelated with business cycles

#3 Wage posted on vacancy may differ from wage for new hire

Solution: compare to new hire wage from **survey** data

# Burning Glass Tracks Wage for New Hires from CPS

Additional



# Job-Level Wages in Burning Glass

# Job-Level Wage for New Hires: Concept

Key advantage of Burning Glass: **job-level** wage for new hires

- Job Level = successive vacancies posted for same job title + establishment

Benefits of job level data:

**#1 Job-level wage** key for unemployment fluctuations in standard model  
(e.g. Gertler & Trigari 2009, Gertler et al 2019)

- Analytical derivation in DMP model with heterogeneous jobs (cf. Elsby & Michaels 2013)

[Details](#)

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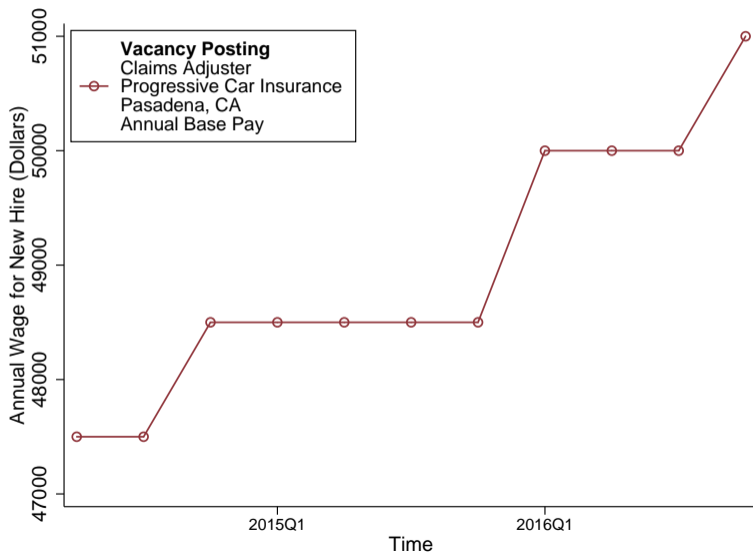
[Details](#)

**#2 Purge wage changes due to job composition**

- Regression of log wages on job fixed effect has R squared > 90%
- Our measure of job captures relevant heterogeneity

# Example: Job-Level Wages

Measurement Details





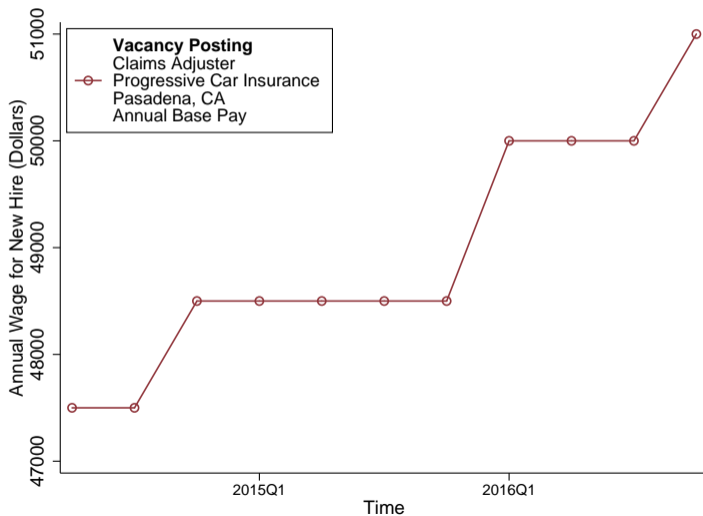
# Wage for New Hires: Rigid Downward and Flexible Upward

# Downward Rigidity for New Hires: Overview

Wage for new hires is rigid downward and flexible upward:

1. **Job-level wages rarely change, rise more often than fall**
2. Job-level wages respond to expansions, do not respond to contractions
3. No downward rigidity in average wage due to job composition
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# Job-level Wages Rarely Change, Rise More Often Than Fall



# Wage for New Hires Changes Infrequently at Job Level

	No weight	Occ. weight	State weight
Duration Unchanged Wages	3.33		
Number of observations	5 667 270		

Annual

# Wage for New Hires Changes Infrequently at Job Level

	No weight	Occ. weight	State weight
Duration Unchanged Wages	3.33	3.03	3.42
Number of observations	5 667 270	5 431 959	5 667 270

Annual

# Job-Level Wages Rise More Often Than Fall

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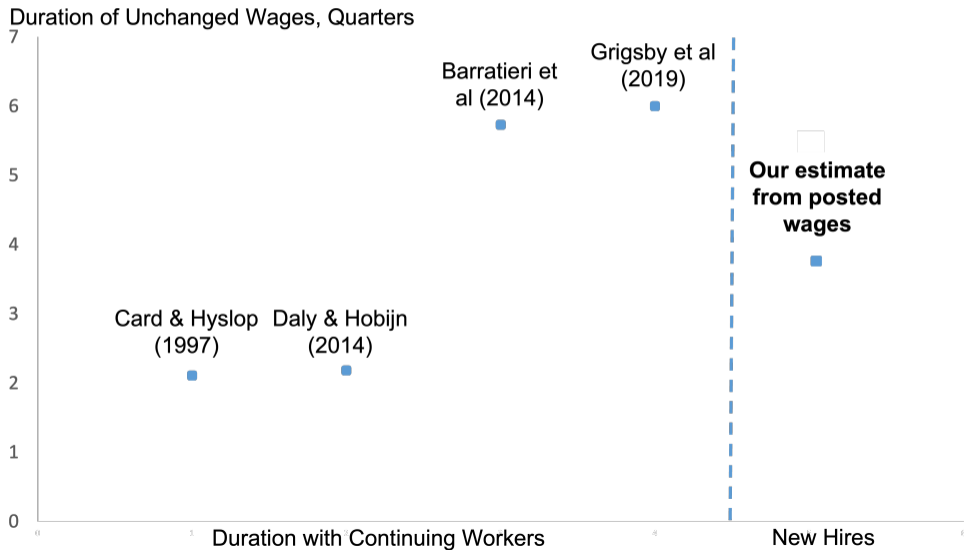


McCrary Test: p value = 0

# Wage Setting For New Hires vs. Continuing Workers



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# Job-Level Wage Changes and Unemployment Changes

- Study response of job-level wages to **state** business cycles
  - States are natural labor market (Yagan 2019; Beraja, Hurst & Ospina 2019) [Details](#)
- Specification

$$\Delta \log w_{ist} = \alpha + \gamma_t + \beta \Delta U_{st} + \delta I[\Delta U_{st} < 0] \Delta U_{st} + \varepsilon_{ist}$$

- $w_{ist}$  = nominal **posted wage**, job  $i$  and quarter  $t$
- $U_{st}$  = quarter-state unemployment, for 2010-2021
- Difference wages at **job level**
- Instrument for  $\Delta U_{st}$  with QCEW employment growth

# Wages Respond to Expansions But Not Contractions



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	$\Delta \log w_{ist}$
$\Delta U_{st}$	0.08 (0.02)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	-0.81 (0.14)
Time Effect	N
$N$	5 554 157

SEs clustered by state

# Wages Respond to Expansions But Not Contractions

	$\Delta \log w_{ist}$	
$\Delta U_{st}$	0.08 (0.02)	0.11 (0.04)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	-0.81 (0.14)	-1.50 (0.25)
Time Effect	N	Y
$N$	5 554 157	5 554 157

SEs clustered by state

# Wages Respond to Expansions But Not Contractions

	$\Delta \log w_{ist}$			
$\Delta U_{st}$	0.08 (0.02)	0.11 (0.04)	0.15 (0.04)	-0.62 (0.11)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	-0.81 (0.14)	-1.50 (0.25)	-1.53 (0.25)	
Time Effect	N	Y	Y	Y
State Trend	N	N	Y	N
$N$	5 554 157	5 554 157	5 554 157	5 545 577

SEs clustered by state

# Wages Respond to Expansions But Not Contractions

	$\Delta \log w_{ist}$				
$\Delta U_{st}$	0.08 (0.02)	0.11 (0.04)	0.15 (0.04)	-0.62 (0.11)	0.00 (0.09)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	-0.81 (0.14)	-1.50 (0.25)	-1.53 (0.25)		-1.28 (0.26)
Time Effect	N	Y	Y	Y	Y
State Trend	N	N	Y	N	N
Real Wages	N	N	N	N	Y
$N$	5 554 157	5 554 157	5 554 157	5 545 577	5 545 577

SEs clustered by state



# Identifying Labor Demand Shocks

- Now: wage rigidity with respect to identified labor demand shocks [Skip](#)

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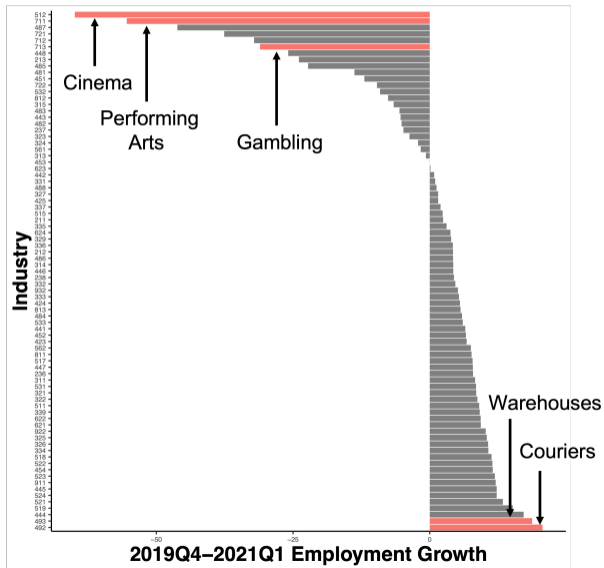
- Now: wage rigidity with respect to identified labor demand shocks Skip

## 1. Industry “shift share” instrument

$$\Delta \text{labor demand}_{st} = \sum_i \text{state industry share}_{is} \times \Delta \text{national industry employment}_{i,-s}$$

ID assumption: variation in industry employment due to labor demand

# Industry Variation Suggests Labor Demand Shocks



# Identifying Labor Demand Shocks

- Now: document downward wage rigidity with respect to labor demand shocks
  - 1. Industry “shift share” instrument  
ID assumption: variation in industry employment due to labor demand
  - 2. Labor supply controls: (e.g. UI replacement rate  $\times$  CARES Act indicator)  
ID assumption: controls absorb all labor supply shocks
  - 3. Oil shock  
ID assumption: regional response to oil shock uncorrelated with labor supply
- (+ All results hold pre-pandemic)

# Identifying Labor Demand

	$\Delta \log w_{ist}$			
	Baseline	Industry Shift Share	Labor Supply Controls	Oil Shock
$\Delta U_{st}$	0.11 (0.04)			
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	-1.50 (0.25)			
Time Effect	Y	Y	Y	Y
$N$	5 554 157	5 554 157	5 504 321	5 552 670

# Identifying Labor Demand

	$\Delta \log w_{ist}$			
	Baseline	Industry Shift Share	Labor Supply Controls	Oil Shock
$\Delta U_{st}$	0.11 (0.04)	0.01 (0.41)	0.06 (0.037)	0.11 (0.21)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	-1.50 (0.25)	-1.52 (0.63)	-1.53 (0.26)	-1.21 (0.43)
Time Effect	Y	Y	Y	Y
$N$	5 554 157	5 554 157	5 504 321	5 552 670

# Downward Rigidity: Robustness and Extensions

- Establishment level results
  - Similar downward rigidity at establishment-level vs. job-level [Details](#)
- Further job-level specifications [Specifications](#)
  - Similar results pre 2020
- Industry evidence (3 digit) [Details](#)
- (Lack of) heterogeneity by occupation [Details](#)
- (Lack of) heterogeneity by degree of wage bargaining [Details](#)
- No evidence of substitution in response to downward wage rigidity [Details](#)
- Downward wage rigidity in calibrated model [Details](#)

# Downward Rigidity for New Hires: Overview

Wage for new hires is rigid downward and flexible upward:

1. Job-level wages rarely change, rise more often than fall
2. Job-level wages rise during expansions, do not fall during contractions
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## Job Composition: Average vs. Job-Level Wages

- Economy:  $I$  job types,  $S$  regions,  $T$  periods, wage for hire  $w_{ist}$ , share  $v_{ist}$
- Wage growth for new hires:

$$\underbrace{\overline{\Delta \log w_{st}}}_{\text{average wage growth}} \approx \underbrace{\sum_i v_{ist} \Delta \log w_{ist}}_{\text{job level wage growth}} + \underbrace{\sum_i \log w_{ist} \Delta v_{ist}}_{\text{wage growth from job composition}}$$

- Prior work: **average** wage growth from **survey data**  
(e.g. Haefke, Sonntag & van Rens 2013)
- Our paper: **job level** wage growth

# Job Composition Raises Variance of Average Wages

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- Variance of average wage growth higher than job-level wage growth
- If composition and job-level wages independent

## Regressions with Average Wages Less Precise

- Our benchmark regression:

$$\Delta \log w_{ist} = \delta_{\text{Job Level}} I[\Delta U_{st} < 0] \Delta U_{st} + \text{controls}_{st} + \text{error}_{st}$$

- Same regression with average wage for new hires:

$$\Delta \overline{\log w}_{st} = \delta_{\text{Average}} I[\Delta U_{st} < 0] \Delta U_{st} + \text{controls}_{st} + \text{error}_{st}$$

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*For finite states and time periods:*

$$\text{standard deviation} \left[ \hat{\delta}_{\text{Average}} | \Delta U_{st} \right] > \text{standard deviation} \left[ \hat{\delta}_{\text{Job Level}} | \Delta U_{st} \right]$$

*if job composition and  $\Delta U_{st}$  are independent*

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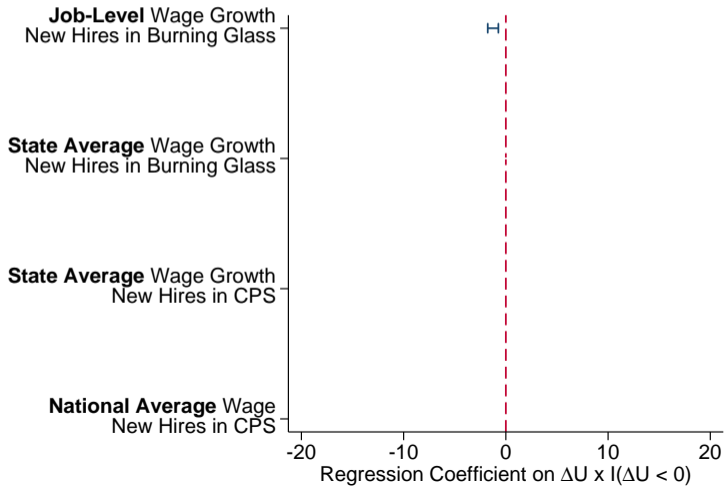
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- Omitted variable bias less important [Details](#)

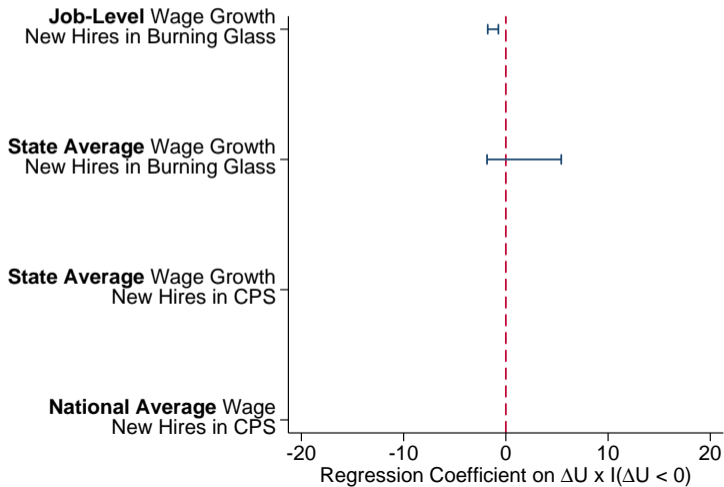
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Regression Outcome Variable



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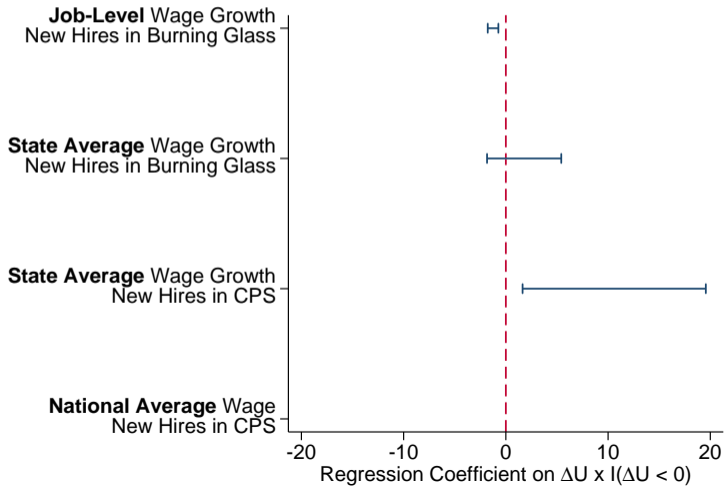
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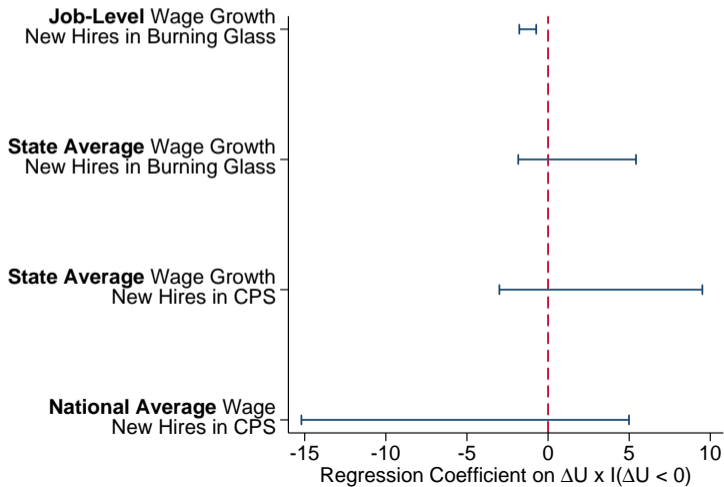
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4. **Downward wage rigidity** → **state dependent wage flexibility upward**

# Downward Rigidity and Wage Flexibility Upward

**Standard model** of downward wage rigidity in DMP model

$$w_t = \max[w_{t-1}, w_t^*] \quad w_t^* = w^*(y_t) \quad \partial w_t^* / \partial y_t > 0$$

where  $y_t$  is revenue product of labor,  $w_t^*$  is Nash bargained wage  
(cf. Schmitt-Grohé & Uribe; Chodorow-Reich & Wieland; Dupraz, Nakamura & Steinsson)

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Model predicts state dependent wage flexibility upward

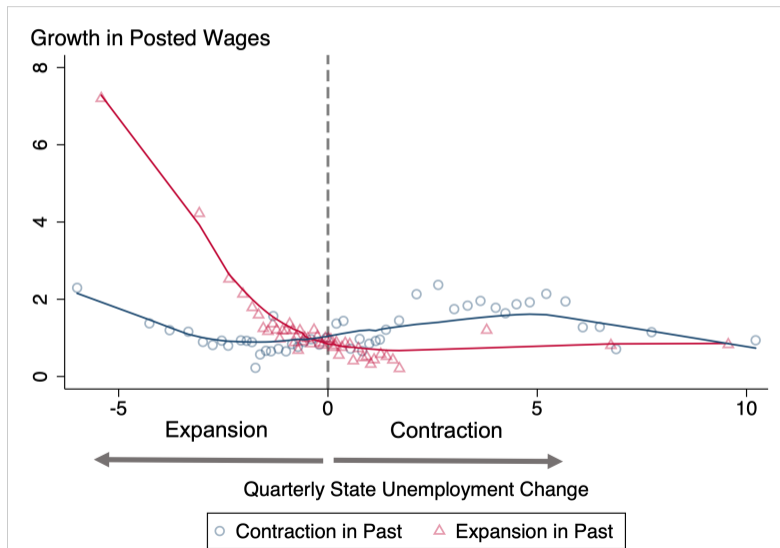
1. Aftermath of large contraction—wages inflexible upward

$$\Delta y_{t-1} \ll 0 \quad \implies \quad w_t = w_{t-1} > w_t^* \quad \implies \quad \partial w_t / \partial_+ y_t = 0$$

2. Aftermath of large expansion—wages flexible upward

$$\Delta y_{t-1} \gg 0 \quad \implies \quad w_t = w_t^* > w_{t-1} \quad \implies \quad \partial w_t / \partial_+ y_t > 0$$

# Wage Flexibility Upward is State Dependent



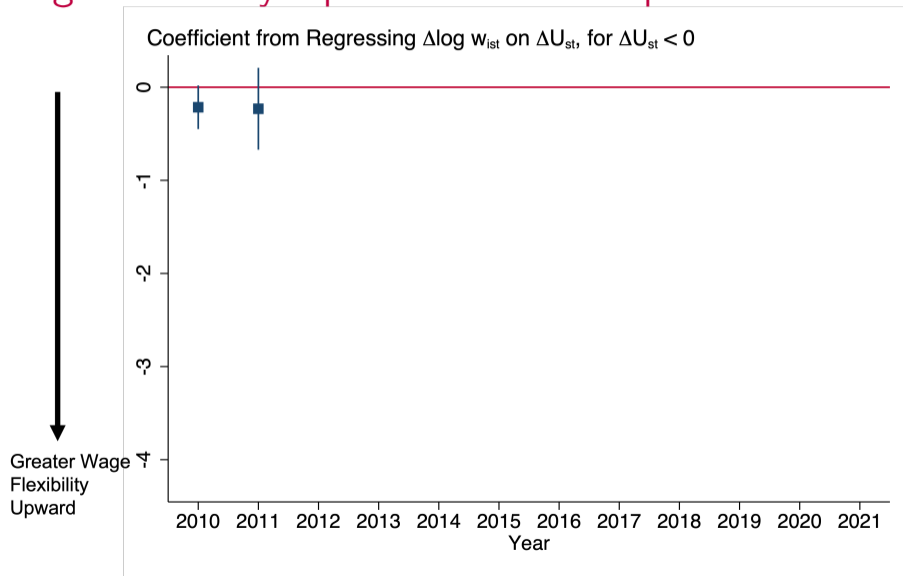
# Wage Flexibility Upward: Regression Estimates

- Regression:

$$\Delta \log w_{ist} = \alpha + \gamma_t + \kappa \Delta U_{st} + \varepsilon_{ist}$$

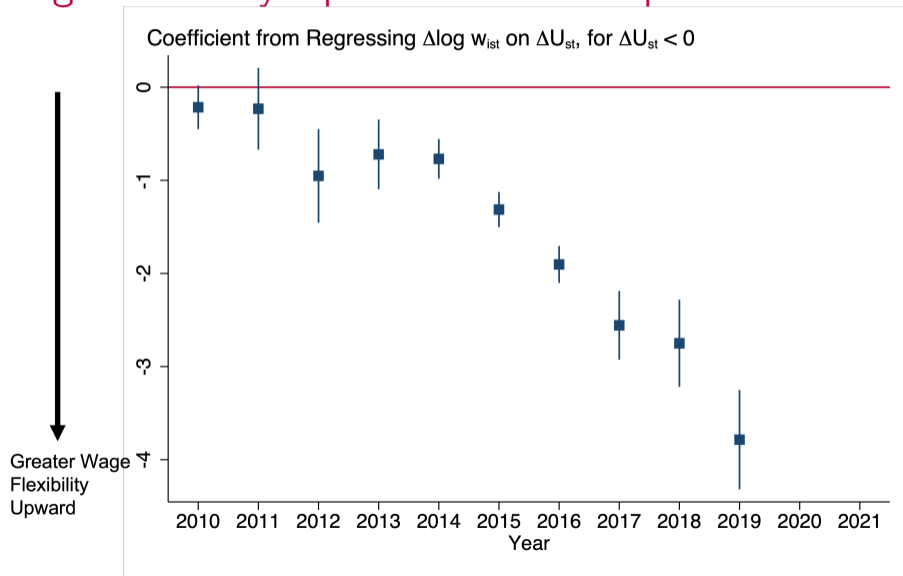
- Study **wage flexibility upward**: restrict sample to  $\Delta U_{st} < 0$
- $\Delta \log w_{ist}$  = growth in wage for new hires at job-level
- $U_{st}$  = quarterly state unemployment
- $\kappa$  is wage flexibility upward
- Estimate regression coefficient  $\kappa_y$  separately for every year

# Wage Flexibility Upward is State Dependent

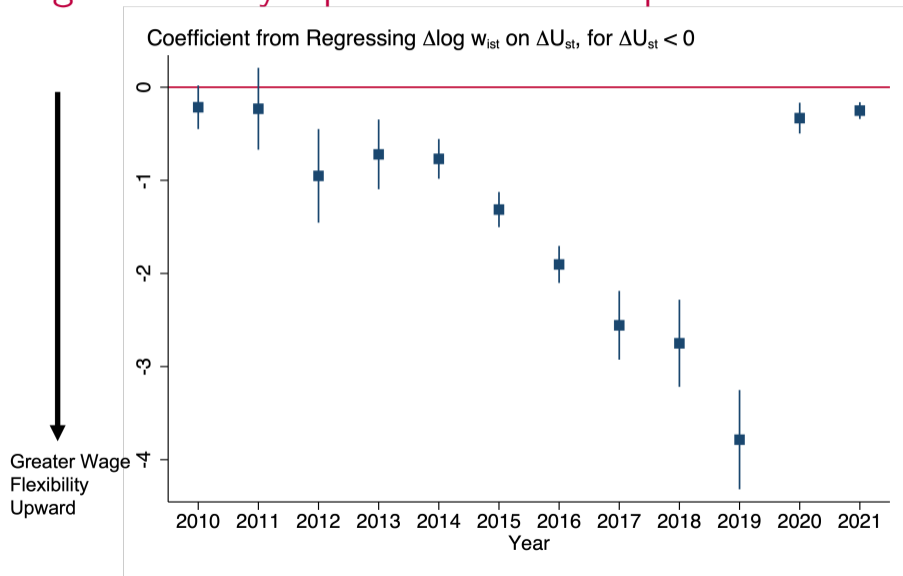




# Wage Flexibility Upward is State Dependent



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# State Dependent Wage Flexibility—Implications

- Estimates of average wage cyclicality hard to interpret
- “Missing wage growth” during 2010-2014 after the Great Recession
- Wage growth may accelerate further after pandemic recession

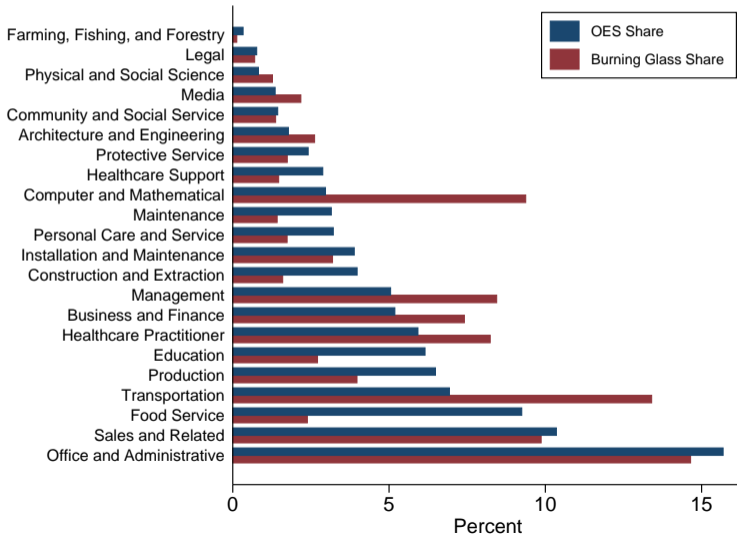
# Conclusion

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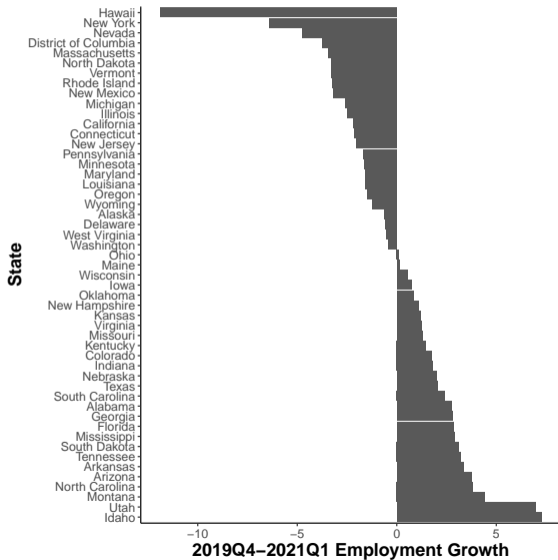
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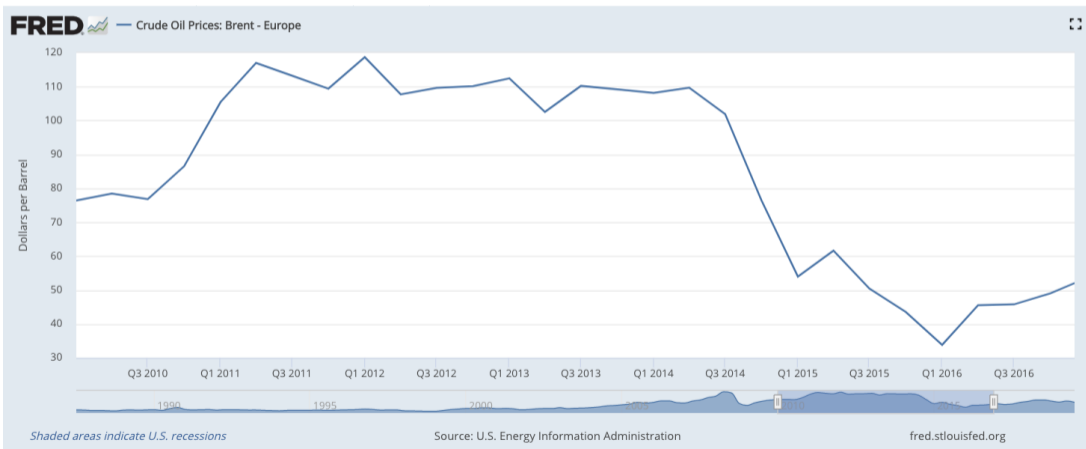
# Occupation Shares in Burning Glass [Return](#)



# Regional Business Cycles [Return](#)



# Oil Price Return





## Further Job-Level Specifications (1/2) [Return](#)

	Coefficient $\Delta U_{st} \times I(\Delta U_{st} < 0)$	S.E.	<i>N</i>
Baseline	-1.501	(0.250)	5 554 157
Occupation weight	-1.617	(0.288)	5 324 569
Control for fill rate	-1.437	(0.292)	3 605 634
No bonuses	-1.760	(0.304)	5 344 277
Region weight	-1.618	(0.275)	5 554 157
Seasonal dummy	-1.854	(0.298)	5 554 157
Seasonal (X-11)	-1.923	(0.310)	5 554 157

## Further Job-Level Specifications (2/2)

	Coefficient $\Delta U_{st} \times I(\Delta U_{st} < 0)$	S.E.	<i>N</i>
4 quarter diff. only	-1.902	(1.213)	346 077
No wage range	-1.824	(0.321)	2 222 025
No time FE	-0.896	(0.174)	5 554 157
No consecutive quarters	-1.734	(0.288)	2 657 855
Pool across pay category	-1.880	(0.311)	4 406 565
Before 2020	-3.175	(0.301)	3 691 677
Oil shock before 2020	-8.905	(2.347)	3 690 214

# Industry Evidence Return

	Quarterly Job-Level Growth in Wage for New Hires			
	(1)	(2)	(3)	(4)
$\Delta \log(\text{employment}_{it})$	-0.0175	-0.0180	-0.0153	-0.0143
	(0.00262)	(0.00314)	(0.00260)	(0.00278)
$\Delta \log(\text{employment}_{it})$	0.0453	0.0480	0.0421	0.0472
$\times I(\Delta \log(\text{employment}_{it}) > 0)$	(0.00916)	(0.00858)	(0.00727)	(0.00789)
Time Effects	Y	Y	Y	Y
Industry Trend	N	Y	N	N
Seasonally Adjusted	N	N	N	Y
Number of observations	2 577 742	2 577 742	2 577 742	2 577 742
Industry clusters	78	78	78	78

## Establishment Robustness (1/2) [Return](#)

	Growth in Wage for New Hires	
	Establishment-Level	Job-Level
$\Delta U_{st}$	0.00392 (0.313)	-0.0517 (0.256)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	-1.082** (0.382)	-1.255*** (0.265)
Time Effect	Y	Y
$N$	1845695	1845695
State Clusters	52	52

## Establishment Robustness (1/2)

Dependent Variable:	Quarterly Establishment Growth in Wage for New Hires			
	(1)	(2)	(3)	(4)
Independent Variable:				
$\Delta U_{st}$	0.00392 (0.313)	-0.268 (0.353)	-0.0431 (0.341)	-0.909*** (0.0737)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	-1.082** (0.382)	-0.785+ (0.427)	-1.021* (0.414)	
Time Effect	Y	Y	Y	Y
State Trend	N	Y	N	N
QCEW Weight	N	N	Y	N
$N$	1845695	1845695	1845695	1845695
State Clusters	52	52	52	52

# Lack of Occupation Heterogeneity Return

Dependent Variable:	Quarterly Job-Level Growth in Wage for New Hires				
Occupation Group:	Management	Services	Sales	Construction	Production
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	-1.177** (0.348)	-1.410*** (0.310)	-0.983* (0.447)	-1.043* (0.433)	-1.552*** (0.321)
Number of Observations	568307	195274	342738	75637	329647

## Job Composition and Average Wages [Return](#)

Dependent Variable:	Quarterly Growth in Wage for New Hires			
	State, CPS		National, CPS	National, NLSY
	(3)	(4)	(5)	(6)
Independent Variable:				
$\Delta$ Unemployment	-5.748 (4.359)	-8.141 (5.903)	3.770 (3.468)	-1.779 (3.172)
$\Delta$ Unemployment $\times$ $I(\Delta$ Unemployment $< 0)$	10.59* (4.560)	13.78 (6.886)	-5.108 (5.151)	2.935 (4.311)
Hagedorn/Manovskii	N	N	N	Y
Cumulative Tightness Control				
<i>N</i>	1377	1377	83	83
State Clusters	51	51	-	-

## Job-Level Wages: Measurement (1/2) [Return](#)

- Define a job as a job title by establishment by pay category
- Restrict to jobs with multiple vacancies
- Take mean posted wage within each job-quarter
- ~5 million vacancies remaining
- Covers 99% of 6 digit occupations by national employment share



## Job-Level Wages: Measurement (2/2)

	Min	Max	Average	Total
Total Vacancy Posts				5 554 157
Share of 6 digit SOC occupations covered in the OES				.99
Posts Per Job	2	23	2.5	
Jobs per 6 digit SOC occupation	1	176081	1247.2	
Jobs per State	264	118 076	19 909	
Jobs per Quarter	7 519	117 566	38 343	

## Annual Probability of Change [Return](#)

	Unweighted	OES Weights	QCEW Weights
Probability of Job-Level Wage Change	0.405	0.418	0.402
Probability of Job-Level Wage Decrease	0.088	0.095	0.09
Probability of Job-Level Wage Increase	0.304	0.305	0.3
Implied Duration Wages Are Unchanged (Years)	1.841	1.836	1.875

# Job-Level Wages Rise More Often Than Fall [Return](#)

	Unweighted	OES Weights	QCEW Weights	High Wage
Prob. of Job-Level Wage Increase	0.11			
Prob. Job-Level Wage Decrease	0.04			

[Annual](#)

## Job-Level Wages Rise More Often Than Fall [Return](#)

	Unweighted	OES Weights	QCEW Weights	High Wage
Prob. of Job-Level Wage Increase	0.11	0.11	0.11	0.11
Prob. Job-Level Wage Decrease	0.04	0.04	0.04	0.04

[Annual](#)

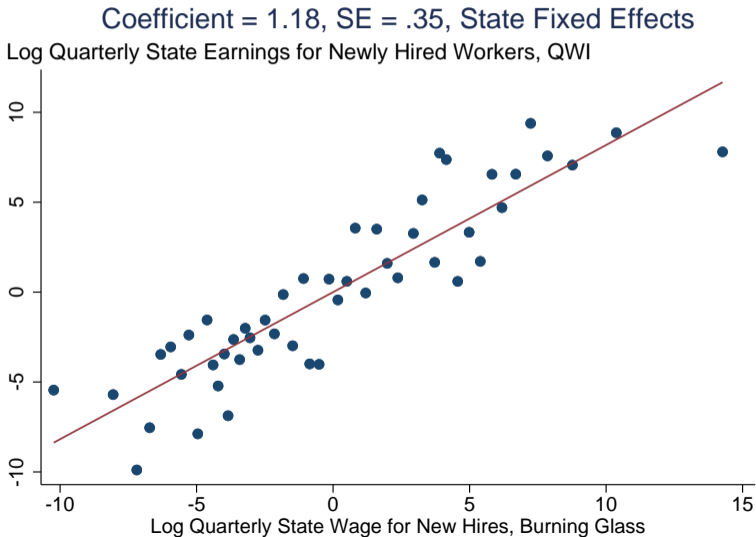
# Composition Bias [Return](#)

Panel A:	Quarterly Change in State Share of High Wage Vacancies			
	(1)	(2)	(3)	(4)
$\Delta U_{st}$	-0.654 (0.831)	-1.040 (1.286)	4.815 (2.677)	-0.0414 (0.393)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	0.982 (1.270)	1.549 (1.927)	-3.537 (5.138)	
Time Effect	Y	Y	Y	Y
State Trend	N	Y	N	N
QCEW Weight	N	N	Y	N
$N$	1404	1404	1404	1404

## Gap Between Burning Glass and CPS Wages [Return](#)

$\Delta [\log(\text{Burning Glass Wage}_{st}) - \log(\text{CPS New Hire Wage}_{st})]$	
$\Delta U_{st}$	-0.0917 (1.431)
State FE	N
N	1377

# Burning Glass Tracks Earnings for New Hires from QWI



# Average Wages: No Rigidity Due to Job Composition

Continue

- Previous work studies **average** wage for new hires
- Our finding:
  - Average wages do not display downward rigidity
  - Even though wages downwardly rigid at job level
  - Key reason is **job composition**
- Due to job composition:
  - Average wages have higher **variance** than job-level wages
  - Regressions with average wages **too imprecise** to detect downward rigidity



# Job-Level Wage Changes Before vs. After 2020 [Return](#)



## State Share of Vacancies Posting Wages [Return](#)

	Change in Share of State Vacancies with Wage			
	(1)	(2)	(3)	(4)
Quarterly State	-0.654	-1.040	4.815	-0.0414
Unemployment Change	(0.831)	(1.286)	(2.677)	(0.393)
Annual State	0.982	1.549	-3.537	
Unemployment Change	(1.270)	(1.927)	(5.138)	
Time Effect	Y	Y	Y	Y
State Trend	N	Y	N	N
QCEW Weight	N	N	Y	N
<i>N</i>	1404	1404	1404	1404

# Job-Level Wages in Model with Heterogeneous Jobs Return

DMP model with heterogeneous jobs as in Elsby & Michaels (2013)

- Continuum of firms  $i \in [0, 1]$ 
  - Heterogeneous and idiosyncratic productivity  $x_{it}$
  - Decreasing returns to scale
  - Pay **job-level wage**  $w_{it}$
- All other parts of the model are standard DMP:
  - Frictional labor market with random search,  $u_t$  unemployed workers
  - Exogenous separations
  - Process for aggregate labor productivity  $y_t$
  - Risk neutral hand-to-mouth workers

# Importance of Job-Level Wages

- Object of interest is  $d \log u / d \log y$ 
  - Elasticity of unemployment to aggregate labor productivity
  - At aggregate steady state (Ljungqvist & Sargent 2017)

# Importance of Job-Level Wages

- Object of interest is  $d \log u / d \log y$ 
  - Elasticity of unemployment to aggregate labor productivity
  - At aggregate steady state (Ljungqvist & Sargent 2017)

*To a first order in a neighborhood of the deterministic steady state, we have*

$$\frac{d \log u}{d \log y} = -A + B \int_0^1 \frac{dw_{it}}{dy} di$$

*for constants  $A, B > 0$ .*

## Lack of Establishment Level Substitution [Return](#)

	Quarterly Change in Share of Establishment Vacancies			
	in High Wage Occupations		with High Wages	
$\Delta U_{st}$	0.158 (0.370)	0.922 (0.642)	-0.0296 (0.0358)	-0.0031 (0.0577)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	0.167 (0.441)	-0.158 (0.735)	0.0537 (0.0369)	0.0173 (.0648)
Time Effect	Y	Y	Y	Y
Size Weighted	N	Y	N	Y
$N$	1770257	1770257	1883361	1883361

## Lack of Market Level Substitution Return

	Quarterly Change in State Share of High Wage Vacancies			
	(1)	(2)	(3)	(4)
$\Delta U_{st}$	-0.654 (0.831)	-1.040 (1.286)	4.815 (2.677)	-0.0414 (0.393)
$\Delta U_{st} \times I(\Delta U_{st} < 0)$	0.982 (1.270)	1.549 (1.927)	-3.537 (5.138)	
State Difference	Y	Y	Y	Y
Time Effect	Y	Y	Y	Y
State Trend	N	Y	N	N
QCEW Weight	N	N	Y	N
<i>N</i>	1404	1404	1404	1404

## Calibrated DMP Model [Return](#)

- Standard DMP model with “reduced form wage rule” (cf. Michaillat 2012)

$$w_t = \max [w_{t-1}, \phi y_t^\gamma]$$

- Calibrate parameters of reduced form wage rule to match our estimates
- Calibrate other parameters to standard values (e.g. Shimer 2005)



# Calibrated DMP Model

	$\Delta \log u_t / \Delta \log y_t$
<b>Values from calibrated model</b>	
Labor demand falling, $\Delta \log y < 0$	-3.00
Labor demand rising, $\Delta \log y > 0$	-0.83
Average value	-1.92
<b>Value from time series data</b>	-1.90

Source: BEA and BLS